



ANSWER BOOK

# Estimators, standard error and confidence intervals

Further Statistics 2 · Year FM

6 questions · 22 marks

NAVY EDITION

## SECTION A · Warm-up

- A1** An estimator is unbiased if its expected value equals the parameter it estimates, so that it is right on average over repeated samples. The sample variance with divisor  $n - 1$ ,  $s^2 = \Sigma(x - \text{sample mean})^2 / (n - 1)$ , is unbiased for  $\sigma^2$ ; dividing by  $n$  instead gives an estimate that is too small on average. BI for the definition, BI for naming  $s^2$  with divisor  $n - 1$ . [2]
- A2** Standard error =  $12/\sqrt{36} = 2$ . The interval is  $48 \pm 1.96 \times 2 = 48 \pm 3.92$ , that is **(44.08, 51.92)**. BI for the standard error, MI for  $48 \pm 1.96 \times 2$ , AI for the interval. [3]

## SECTION B · Standard

- B1** Standard error =  $12/\sqrt{36} = 2$ , unchanged, since the confidence level affects only the multiplier. [4]  
Using  $z = 2.5758$ :  $48 \pm 2.5758 \times 2 = 48 \pm 5.15$ , that is **(42.85, 53.15)**.  
It is wider than the 95% interval by about 1.23 units at each end. More confidence is bought with less precision, and only a larger sample improves both at once. BI for  $z = 2.5758$ , MI for  $48 \pm z \times 2$ , AI for the interval, BI for the comparison. Quoting 2.58 rather than 2.5758 is acceptable, but 1.96 for a 99% interval loses every accuracy mark that follows.
- B2** The total width is  $2 \times 1.96 \times 12/\sqrt{n}$ , and the requirement is that this is at most 4. [4]  
Working with the total width rather than the half-width is where most answers go wrong, and it changes the answer by a factor of four.  
Rearranging:  $\sqrt{n} \geq 2 \times 1.96 \times 12/4 = 11.76$ , so  $n \geq 138.3$ .  
Since  $n$  must be a whole number and the inequality runs upwards, round **up** to  **$n = 139$** .  
Rounding down to 138 gives an interval slightly too wide and loses the final mark. MI for an inequality in the total width, AI for  $\sqrt{n} \geq 11.76$ , dMI for reaching  $n \geq 138.3$ , AI for 139.  
Halving the width again would need four times as many readings, since  $n$  appears under a square root.
- B3** The population mean is a fixed number rather than a random one, so once the interval has been calculated it either contains  $\mu$  or it does not, and no probability remains. What varies from sample to sample is the interval, not  $\mu$ . [3]  
The correct statement is that the method produces intervals of which 95% would contain  $\mu$  in repeated sampling. The confidence attaches to the procedure, not to the one interval in front of you. BI for  $\mu$  being a fixed number, BI for the interval being what varies, BI for the repeated-sampling statement.

## SECTION C · Stretch

- C1 Unbiasedness.**  $E(\bar{X}) = E(\bar{Y}) = \mu$ , so  $E(T) = \lambda\mu + (1 - \lambda)\mu = \mu$  for every  $\lambda$ . The coefficients were chosen to add to 1, which is exactly what unbiasedness requires, so no value of  $\lambda$  can be ruled out on those grounds. Unbiasedness alone never settles which estimator to use. [6]
- Variance.** The two sample means are independent, with  $\text{Var}(\bar{X}) = \sigma^2/4$  and  $\text{Var}(\bar{Y}) = \sigma^2/6$ . Each coefficient squares:  
 $\text{Var}(T) = \lambda^2\sigma^2/4 + (1 - \lambda)^2\sigma^2/6$ .
- Minimising.** Differentiate with respect to  $\lambda$  and set the result to zero:  $\lambda\sigma^2/2 - (1 - \lambda)\sigma^2/3 = 0$ . The  $\sigma^2$  cancels, which is the sign that the answer cannot depend on it. So  $3\lambda = 2(1 - \lambda)$ , giving  $5\lambda = 2$  and  $\lambda = 0.4$ . The second derivative is positive, so this is a minimum; say so, because it is a mark.
- Then  $\text{Var}(T) = 0.16\sigma^2/4 + 0.36\sigma^2/6 = 0.04\sigma^2 + 0.06\sigma^2 = \mathbf{0.1\sigma^2}$ , that is  $\sigma^2/10$ .
- B1 for  $E(T) = \mu$  for every  $\lambda$ , M1 A1 for  $\text{Var}(T) = \lambda^2\sigma^2/4 + (1 - \lambda)^2\sigma^2/6$ , M1 for differentiating and equating to zero, A1 for  $\lambda = 0.4$ , A1 for  $0.1\sigma^2$ .
- Note what that means.  $\sigma^2/10$  is the variance of the mean of all ten observations, so the best weighting is the one that treats every original reading equally. Weighting the two sample means equally, at  $\lambda = 0.5$ , gives  $5\sigma^2/48 \approx 0.104\sigma^2$ , which is worse.